

CSCE 619-600

Networks and Distributed Processing

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## Renewal Process Theory II

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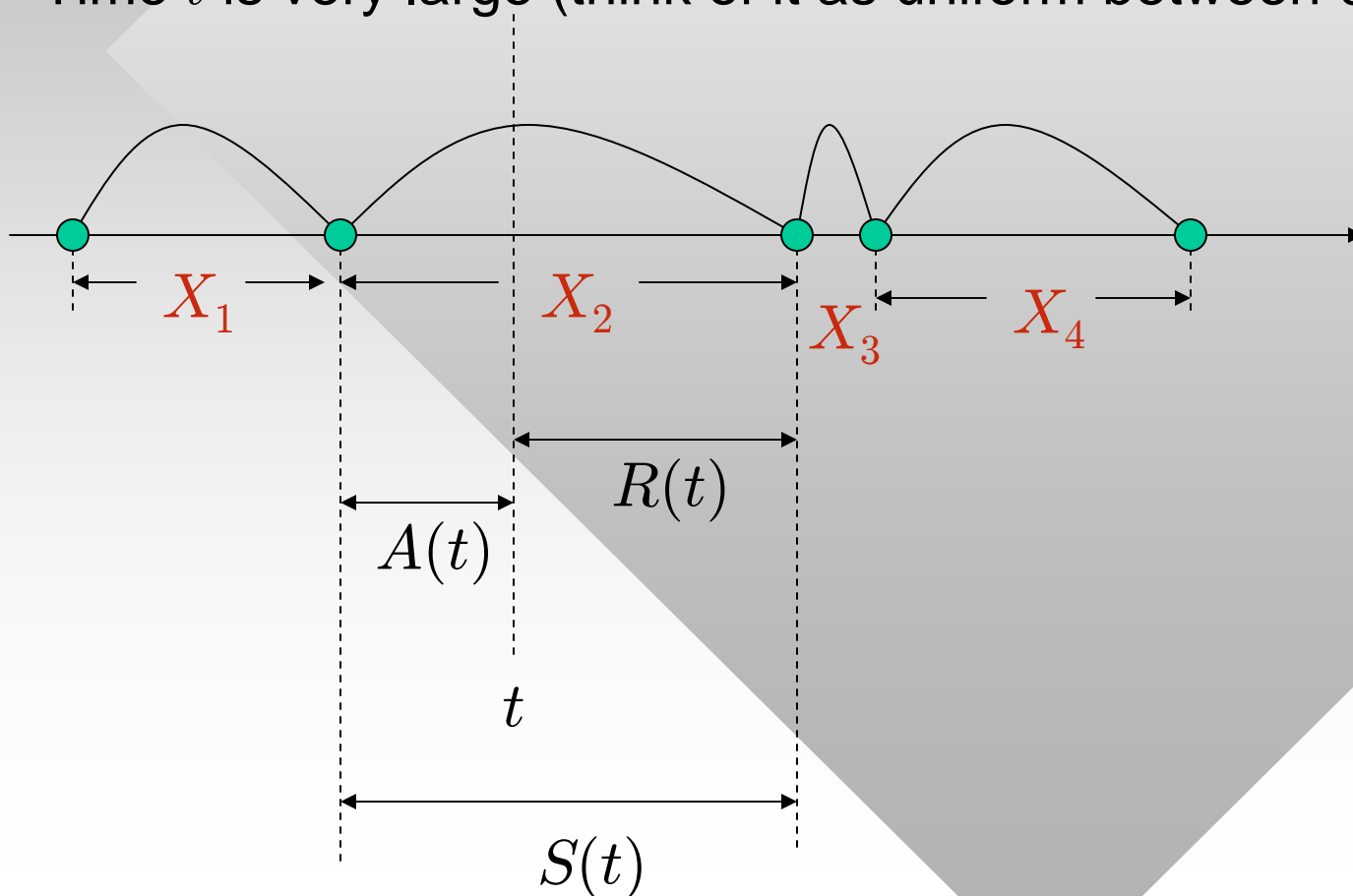
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# Agenda

- Inspection paradox
- Renewal-reward processes
  - Renewal-reward theorem
  - Examples
- Distribution of residuals and age
  - Derivation and examples
- Distribution of spread
  - Examples

# Inspection Paradox

- Recall from the previous lecture
  - Age  $A(t)$ , residual delay  $R(t)$ , and spread  $S(t)$
  - Time  $t$  is very large (think of it as uniform between 0 and  $\infty$ )



# Inspection Paradox 2

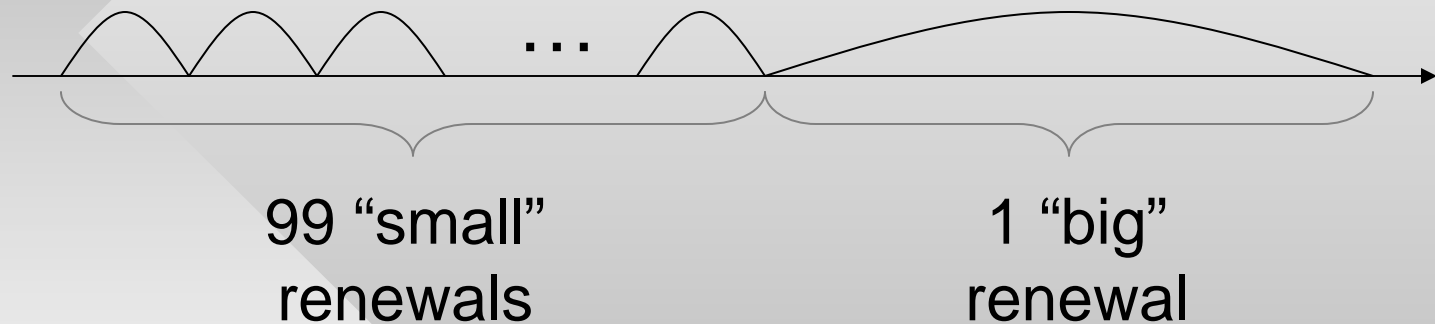
- The “inspection paradox”
  - When we randomly examine a renewal process at time  $t$ , the average spread  $S(t)$  can be much larger than  $E[X_j]$
- This can be shown through an example
  - Consider two types of inter-bus delays:

$$P(X_j = 1) = 0.99, P(X_j = 100) = 0.01$$

- This means that 99% of all renewal intervals are 1 minute and the remaining 1% are 100 minutes
  - What is the probability to observe interval  $X_j = 100$  at some large time instance  $t$ ?
- Clearly, not 1%

# Inspection Paradox 3

- Notice that in the “average sense,” the process has this structure:



- Thus, each 199 minutes (on average) contain one large and 99 small intervals
  - Therefore, the probability that  $t$  randomly “lands” into a large interval is  $100/199 \approx 50\%$  instead of 1%

# Inspection Paradox 4

- A random observer is more likely to inspect the process during large intervals and experience longer wait times than if he/she arrived at renewal points  $Z_n$
- Definition (pg. 22)
  - $X$  is *stochastically larger* than  $Y$  if

$$P(X > t) \geq P(Y > t), \forall t$$

- For example, uniform  $X$  in  $[3,5]$  is stochastically larger than uniform  $Y$  in  $[2,4]$
- Theorem: random variable  $S(t)$  is stochastically larger than  $X_j$

# Inspection Paradox 5

- We now have an intuitive explanation of why Pareto wait times in the bus example became larger
- This is a consequence of very large samples  $X_j$  in the Pareto distribution during which we are more likely to inspect the system
- Example:
  - Compute the expected wait time for the example shown two slides back
  - Recall:

$$E[X] = \sum_y E[X|Y = y]P(Y = y)$$

# Inspection Paradox 6

- Thus:

$$E[W] = E[W|small]P(small) + E[W|large]P(large)$$

- where

$$P(small) = 99/199, P(large) = 100/199$$

- Putting the pieces together:

$$E[W] = \frac{1}{2} \times \frac{99}{199} + \frac{100}{2} \times \frac{100}{199} = 25.37...$$

- Also note that  $E[X_j]$  is only 1.99 minutes
  - This means that buses arrive at the rate of 30 per hour; however, your expected wait is over 25 minutes!



# Renewal Rewards

- Our goal today is to derive the distribution of residual waiting time  $R(t)$  and spread  $S(t)$ 
  - Before we do that, we need several results from renewal rewards, which is a segment of renewal process theory
- Suppose each renewal  $j$  earns reward  $P_j$ 
  - Sequence  $\{P_1, P_2, \dots\}$  consists of iid random variables
- Then, the cumulative reward in  $[0, t]$  is

$$C(t) = \sum_{j=1}^{M(t)} P_j$$

- Technical note
  - Rewards  $P_j$  may depend on  $X_j$ , but not  $X_i$ ,  $i \neq j$

# Renewal Rewards 2

- Let the *expected reward* be  $c(t) = E[C(t)]$
- The Renewal-Reward Theorem:

$$\lim_{t \rightarrow \infty} \frac{C(t)}{t} = \lim_{t \rightarrow \infty} \frac{c(t)}{t} = \frac{E[P_j]}{E[X_j]} = \mu E[P_j]$$

- Example:
  - You are taking a long test that lasts  $t$  time units
  - Suppose it takes  $X_j$  seconds to solve problem  $j$
  - You earn  $P_j$  points for solving the  $j$ -th problem
  - What is your score at the end of the test?
  - Approximately

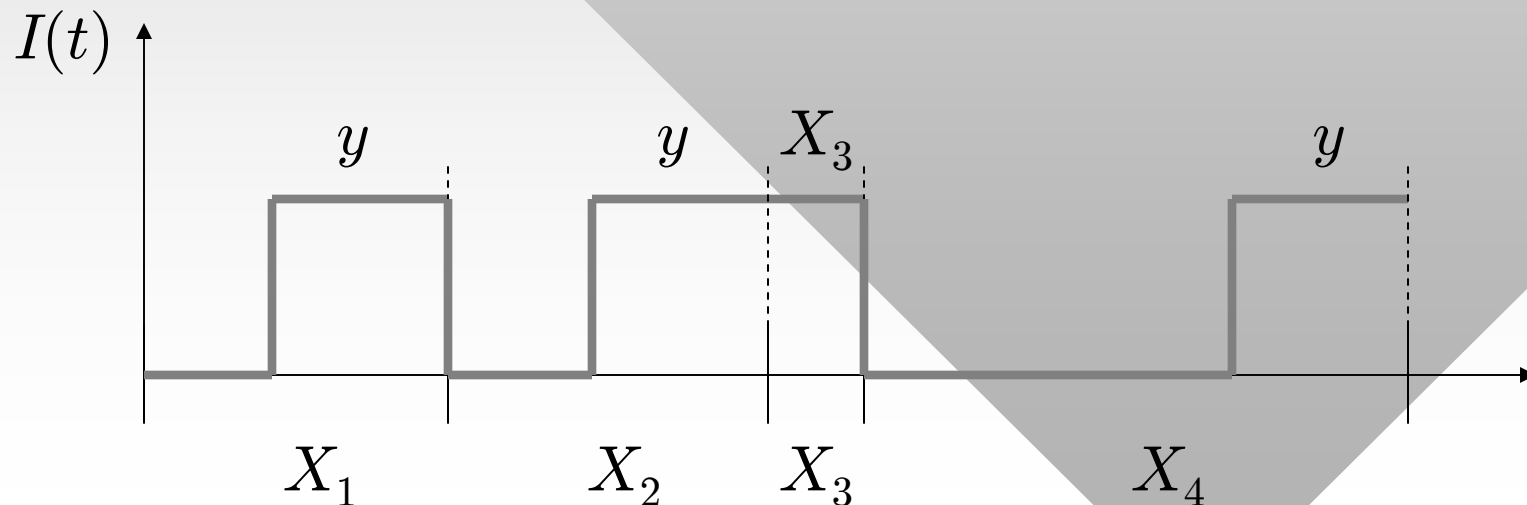
$$t \frac{E[P_j]}{E[X_j]}$$

# Distribution of Residuals

- We are ready to derive the distribution of  $R(t)$
- Fix  $y \geq 0$  and define indicator process:

$$I(t) = \begin{cases} 1 & R(t) \leq y \\ 0 & \text{otherwise} \end{cases}$$

- Graphically, it looks like this:



## Distribution of Residuals 2

- If we throw a random point  $t$  on the interval  $[0, T]$ , what is the probability to hit a segment with  $I(t) = 1$ ?

$$P(R(t) \leq y) = \frac{1}{T} \int_0^T I(u) du$$

- Setting  $R$  as the limit of  $R(t)$  for  $t \rightarrow \infty$

$$P(R \leq y) = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T I(u) du$$

- All we now have to do is compute this limit
  - Define reward  $P_j$  for renewal  $j$  to be:

$$P_j = \int_{Z_{j-1}}^{Z_j} I(u) du = \begin{cases} X_j & X_j \leq y \\ y & X_j > y \end{cases}$$

- Alternatively,  $P_j = \min(X_j, y)$

## Distribution of Residuals 3

- Next observe that the integral of  $I(u)$  is actually the cumulative reward at time  $T$ :

$$C(T) = \sum_{j=1}^{M(T)} P_j = \int_0^T I(u) du$$

- We neglect special cases when  $T$  does not fall on the boundary of  $Z_j$  (i.e, partial rewards) as their contribution tends to zero for large  $T$
- Next, apply the renewal reward theorem to  $C(T)$ :

$$\lim_{T \rightarrow \infty} \frac{C(T)}{T} = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T I(u) du = \frac{E[P_j]}{E[X_j]}$$

# Distribution of Residuals 4

- Observe that the middle term in the last equation is  $P(R \leq y)$  and re-write:

$$P(R \leq y) = \frac{E[P_j]}{E[X_j]}$$

- What is left is to obtain  $E[P_j]$
- Lemma: Suppose  $X$  is a non-negative variable and  $a > 0$  is some constant. Then

$$\begin{aligned} E[\min(X, a)] &= \int_0^{\infty} P(\min(X, a) > x) dx \\ &= \int_0^a P(\min(X, a) > x) dx \\ &= \int_0^a P(X > x) dx \end{aligned}$$

- Reason: for any  $x \leq a$ , it follows that  $X > x$  iff  $\min(X, a) > x$

# Distribution of Residuals 5

- Combining the pieces:

$$P(R \leq y) = \frac{1}{E[X_j]} \int_0^y P(X_j > x) dx$$

- Or in a more digestible form:

$$P(R \leq y) = \frac{1}{E[X_j]} \int_0^y (1 - F(x)) dx$$

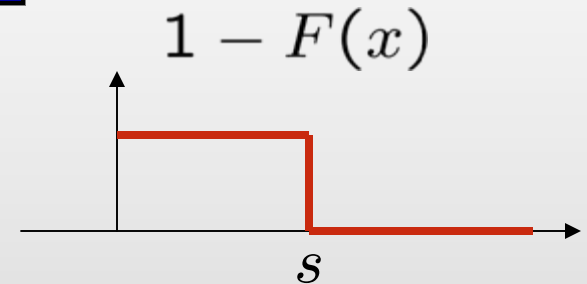
- Note: the same technique applies to age:

$$P(A \leq y) = \frac{1}{E[X_j]} \int_0^y (1 - F(x)) dx$$

# Distribution of Residuals 6

- Examples:

- Constant inter-bus delay  $s$ :



$$P(R \leq y) = \begin{cases} y/s & y \leq s \\ 1 & y > s \end{cases} \quad \leftarrow \text{uniform}$$

- Exponential:

$$P(R \leq y) = \lambda \int_0^y e^{-\lambda x} dx = 1 - e^{-\lambda y}$$

exponential

- Pareto:

$$P(R \leq y) = \frac{\alpha - 1}{\beta} \int_0^y (1 + x/\beta)^{-\alpha} dx = 1 - (1 + y/\beta)^{1-\alpha}$$

Pareto



# Distribution of Residuals 7

- This explains the observations in homework #1
  - Inter-bus delays with Pareto  $\alpha = 3$  had wait times that were Pareto with  $\alpha = 2$
  - It also confirms that when the delay between buses is constant, your expected wait is uniform in  $[0, s]$
- The final important result to derive is the expected wait time  $E[R]$
- Define the residual CDF (also called the *equilibrium* distribution) of  $F$ :

$$F_R(y) := P(R \leq y) = \frac{1}{E[X_j]} \int_0^y (1 - F(x)) dx$$

# Distribution of Residuals 8

- Then the density of wait time is:

$$f_R(y) = F'_R(y) = \frac{1 - F(y)}{E[X_j]}$$

- Next, we compute the expected wait time:

$$E[R] = \int_0^{\infty} y f_R(y) dy = \frac{1}{E[X_j]} \int_0^{\infty} y(1 - F(y)) dy$$

- To solve the integral, notice:

$$E[X^2] = \int_0^{\infty} P(X^2 > x) dx = \int_0^{\infty} P(X > x^{1/2}) dx$$

# Distribution of Residuals 9

- Substituting  $u = x^{1/2}$  ( $dx = 2udu$ ):

$$E[X^2] = \int_0^{\infty} 2uP(X > u)du = 2 \int_0^{\infty} u(1 - F(u))du$$

- Recalling the equation on  $E[R]$ :

$$E[R] = \frac{1}{E[X_j]} \int_0^{\infty} y(1 - F(y))dy$$

- We have:

$$E[R] = \frac{E[X_j^2]}{2E[X_j]}$$

# Distribution of Residuals 10

- Examples:
  - Constant inter-bus delay:

$$E[R] = \frac{s^2}{2s} = \frac{s}{2}$$

- Exponential:

$$E[X^2] = Var[X] + E^2[X] = 2/\lambda^2 \Rightarrow$$

$$E[R] = \frac{2/\lambda^2}{2/\lambda} = 1/\lambda = E[X_j]$$

- Pareto: derive at home

# Distribution of Spread

- Using similar techniques, we can derive the limiting distribution of spread  $S(t)$
- The only difference is that we define

$$P_j = \begin{cases} X_j & X_j \leq y \\ 0 & X_j > y \end{cases}$$

- The rest of the derivations are similar and lead to:

$$F_S(y) := P(S \leq y) = \frac{1}{E[X_j]} \left( yF(y) - \int_0^y F(x)dx \right)$$

- and its density is:

$$f_S(y) = \frac{yf(y)}{E[X_j]}$$

# Wrap-up

- Example

- Spread distribution for exponential  $X_j$ :

$$f_S(y) = \frac{yf(y)}{E[X_j]} = \frac{y\lambda e^{-\lambda y}}{1/\lambda} = \lambda^2 y e^{-\lambda y}$$

- Interestingly, this is the same density as that of a sum of two exponential random variables, i.e., Erlang(2)

- Why was this intuitively expected?

- For exponential  $X_j$ , average spread  $E[S] = 2E[X_j]$

- In general,

$$E[S] = E[R] + E[A] = 2E[R] = \frac{E[X_j^2]}{E[X_j]}$$